

# Exercise Session: MCMC methods

## Exercise 1 (Simulation of a Pair)

Let  $(X, Y)$  be a random pair with joint density

$$f(x, y) = e^{-y} \mathbf{1}_{0 \leq x \leq y}.$$

1. Determine the marginal density of  $X$ , denoted  $f(x)$ . What distribution do you recognize?
2. Given  $X = x \geq 0$ , determine the conditional density  $f(y|x)$ . What distribution do you recognize?
3. Deduce a method to simulate a realization of the random pair  $(X, Y)$ .
4. Given  $Y = y \geq 0$ , determine the conditional density  $f(x|y)$ . What distribution do you recognize?
5. Starting, for example, from the point  $(x_0, y_0) = (0, 1)$ , write the pseudo-code of a Gibbs sampler with deterministic scanning to obtain a sample  $(X_1, Y_1) \dots, (X_n, Y_n)$  to sample from the target law  $f$ .
6. Between the two proposed methods, which one would you choose to simulate according to the density  $f(x, y)$ ?

## Exercise 2 (Metropolis-Hastings Algorithm)

Let  $E$  be a finite state space and  $\pi$  a probability measure on  $E$  such that  $\pi(x) > 0$  for any  $x \in E$ . We recall the Metropolis-Hastings algorithm for an auxiliary transition kernel  $Q$  (such that  $Q(x, y) = 0$  if and only if  $Q(y, x) = 0$ )

Step 0:

Initialize  $X_0$  ;

Step n+1:

Choose  $y$  according to the law  $Q(X_n, \cdot)$  ;

Set  $X_{n+1} = y$  with probability  $\min\left(1, \frac{\pi(y)Q(y, X_n)}{\pi(X_n)Q(X_n, y)}\right)$ , otherwise set  $X_{n+1} = X_n$

- 1) Show that if the acceptance-rejection probability  $\min\left(1, \frac{\pi(y)Q(y, x)}{\pi(x)Q(x, y)}\right)$  is replaced by

$$\frac{\pi(y)Q(y, x)}{\pi(y)Q(y, x) + \pi(x)Q(x, y)},$$

the measure  $\pi$  will still be invariant for the Markov chain defined by the Metropolis-Hastings algorithm.

- 2) More generally, replacing the acceptance-rejection probability  $\min\left(1, \frac{\pi(y)Q(y, x)}{\pi(x)Q(x, y)}\right)$  by

$$\alpha\left(\frac{\pi(y)Q(y, x)}{\pi(x)Q(x, y)}\right),$$

with  $\alpha : \mathbb{R}_+ \rightarrow (0, 1]$ , give a sufficient condition on the function  $\alpha$  for the measure  $\pi$  to be the invariant measure of the Markov chain defined by the Metropolis-Hastings algorithm.

### Exercise 3 (Joint law and Gibbs sampler)

We consider the density

$$f(x, y) = C \exp\left(-\frac{y^2}{2} - \frac{x^2(y^2 + 4y + 4)}{2}\right).$$

The renormalizing constant  $C > 0$  is not specified here, but it does not matter for this exercise.

1. What is the law of  $X$  conditional on  $Y = y$ ? What is the law of  $Y$  conditional on  $X = x$ ?
2. Propose a deterministic scan Gibbs sampler pseudo-code to sample from the target  $f$ .

### Exercise 4 (Markov chain)

For  $\theta \in [0, \frac{1}{2}]$ , we consider the Markov chain on  $E = \{1, 2, 3\}$  with initial state  $X_0 = 1$  and transition matrix

$$\Pi = \begin{pmatrix} \frac{1}{2} & \frac{\theta}{2} & \frac{1}{2}(1-\theta) \\ \theta & 1-2\theta & \theta \\ 0 & 1 & 0 \end{pmatrix}.$$

1. For what values of  $\theta$  is the chain irreducible?
2. Suppose  $\theta = 0$  and  $X_1 = 1$ . Compute  $\mathbb{P}(X_3 = 1 | X_1 = 1)$ .
3. Let  $k \in \{1, 2, 3\}$ . Suppose  $\theta = \frac{1}{3}$ . What is the almost sure limit of  $\frac{1}{n} \sum_{j=1}^n \mathbf{1}_{X_j=k}$ ?

### Exercise 5 (Simultaneous Gibbs = Flawed Gibbs)

Consider the distribution  $\pi = \pi_{X,Y}$  over  $E = \{(0,0), (0,1), (1,0), (1,1)\}$  defined as follows

$$\pi(0,0) = \frac{2}{5}, \quad \text{and} \quad \pi(0,1) = \pi(1,0) = \pi(1,1) = \frac{1}{5}.$$

On this state space, we consider the Markov chain  $(X_n)$  defined by the following transitions: Letting  $(x, y) = (X_t, Y_t)$ , draw independently

$$X_{t+1} \sim \pi_{X|Y}(\cdot | y) \quad \text{and} \quad Y_{t+1} \sim \pi_{Y|X}(\cdot | x)$$

This way of moving in the state space is thus akin to a Gibbs sampler where all coordinates are updated simultaneously (that is to say, not using the most recent updates available so far).

1. Compute the probability distributions  $\pi_{X|Y}(\cdot | y)$  and  $\pi_{Y|X}(\cdot | x)$  for any  $x, y \in \{0, 1\}$  and show that  $\pi_{X|Y} = \pi_{Y|X}$ . Therefore, we will write

$$\pi_{a,b} := \pi_{X|Y}(a | b) = \pi_{Y|X}(a | b), \quad \forall a, b \in \{0, 1\}.$$

2. Justify that, for any  $x, x', y, y' \in \{0, 1\}$ , we have

$$\mathbb{P}((X_{t+1}, Y_{t+1}) = (x', y') | (X_t, Y_t) = (x, y)) = \pi_{x'y} \pi_{y'x}.$$

3. Show that the transition matrix  $P$  of the Markov chain  $(X_n)$  is

	(0, 0)	(0, 1)	(1, 0)	(1, 1)
(0, 0)	$\frac{4}{9}$	$\frac{2}{9}$	$\frac{2}{9}$	$\frac{1}{9}$
(0, 1)	$\frac{1}{3}$	$\frac{1}{6}$	$\frac{1}{3}$	$\frac{1}{6}$
(1, 0)	$\frac{1}{3}$	$\frac{1}{3}$	$\frac{1}{6}$	$\frac{1}{6}$
(1, 1)	$\frac{1}{4}$	$\frac{1}{4}$	$\frac{1}{4}$	$\frac{1}{4}$

Is this chain irreducible? Is it aperiodic?

4. Show that the proposed algorithm does not converge to the desired probability density.
5. Show that the transition matrix  $P_a$  for the Gibbs sampler with random scanning is

	$(0, 0)$	$(0, 1)$	$(1, 0)$	$(1, 1)$
$(0, 0)$	$\frac{2}{3}$	$\frac{1}{6}$	$\frac{1}{6}$	$0$
$(0, 1)$	$\frac{1}{3}$	$\frac{5}{12}$	$0$	$\frac{1}{4}$
$(1, 0)$	$\frac{1}{3}$	$0$	$\frac{5}{12}$	$\frac{1}{4}$
$(1, 1)$	$0$	$\frac{1}{4}$	$\frac{1}{4}$	$\frac{1}{2}$

Verify that  $\pi$  is stationary for  $P_a$ . Is the chain reversible?